

# Uncertain Productivity Growth and the Choice between FDI and Export

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## **WEB-APPENDIX**

This document provides analytical derivations referred to in the article: *Uncertain Productivity Growth and the Choice between FDI and Export*. Published in the Review of International Economics.

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# Appendix

## A Equivalent Risk-Adjusted Return

Given the Geometric Brownian motion in equation (22), from Ito's lemma we have:

$$\mathbb{E} \left( \frac{d\phi^\kappa}{\phi^\kappa} \right) = \left( \kappa\phi^{\kappa-1}d\phi + \frac{1}{2}\kappa(\kappa-1)\phi^{\kappa-2}\sigma^2\phi^2 dt \right) / \phi^\kappa.$$

Substituting for  $d\phi$  leads to

$$\mathbb{E} \left( \frac{d\phi^\kappa}{\phi^\kappa} \right) = \left( \alpha\kappa + \frac{1}{2}\kappa\sigma^2(\kappa-1) \right) dt + \kappa\sigma dz_t. \quad (\text{A.1})$$

From the quadratic equation in (B.1), which is valid for  $\phi^\kappa$  with  $\Psi(\kappa) = 0$ , it follows that

$$\frac{1}{2}\kappa\sigma^2(\kappa-1) = r - (r - (\mu - \alpha))\kappa.$$

Hence, the equivalent risk-adjusted rate of return for an exponential variable results as

$$\mu_e = r + \kappa(\mu - r). \quad (\text{A.2})$$

## B The Fundamental Quadratic Equation

Substituting the guess solution  $F(i, \phi(\sigma)) = A_1\phi^\beta$  and its derivatives into the linear differential equation (27), we receive the fundamental quadratic equation

$$\Psi = \frac{1}{2}\sigma^2\beta(\beta-1) + (r - (\mu - \alpha))\beta - r = 0. \quad (\text{B.1})$$

Consider the total differential

$$\frac{\partial\Psi}{\partial\beta} \frac{\partial\beta}{\partial\sigma} + \frac{\partial\Psi}{\partial\sigma} = 0, \quad (\text{B.2})$$

which can be evaluated at  $\beta = \beta_1$ . The quadratic equation (B.1) increases in  $\beta_1$  with  $\partial\Psi/\partial\beta_1 > 0$ . The derivative of  $\Psi$  with respect to  $\sigma$  results as

$$\frac{\partial\Psi}{\partial\sigma} = \sigma\beta_1(\beta_1 - 1) > 0, \quad (\text{B.3})$$

because of (28). From B.2 we have necessarily  $\frac{\partial \beta_1}{\partial \sigma} < 0$ . Furthermore, the discount rate of periodical profits in equation (25) turns out to be the negative expression of  $\Psi$  evaluated at  $\kappa$ . Note that

$$\mu_e - \alpha_e = r - (r - (\mu - \alpha))\kappa - \frac{1}{2}\kappa(\kappa - 1)\sigma^2. \quad (\text{B.4})$$

For bounded results, this discount rate needs to be strictly positive and, hence,  $\kappa$  must lie between the two roots, specifically:  $\beta_1 > \kappa > 0$ . As a consequence,

$$\frac{\partial \left( \frac{\beta_1}{\beta_1 - \kappa} \right)}{\partial \sigma} > 0. \quad (\text{B.5})$$

For  $\sigma = 0$ , we have  $\mu = r$ , and from equation (B.1) it follows that  $\beta_1 = \frac{r}{\alpha} = \beta$ .

## C Expected Entry Time

By using the Girsanov theorem<sup>1</sup> it is possible to derive the probability density function of the waiting time  $T(i)$  as

$$f(T(i, \phi_\sigma^*(i)), \phi) = \frac{\ln \left( \frac{\phi_\sigma^*(i)}{\phi} \right)}{\sqrt{2\pi\sigma^2 T(i, \phi_\sigma^*(i))}^3} e^{-\frac{\left( \ln \left( \frac{\phi_\sigma^*(i)}{\phi} \right) - (\alpha - \frac{1}{2}\sigma^2) T(i, \phi_\sigma^*(i)) \right)^2}{2\sigma^2 T(i, \phi_\sigma^*(i))}} \quad (\text{C.1})$$

with  $\phi_\sigma^*(i) > \phi$ . The Laplace transform of  $T(i, \phi_\sigma^*(i))$  is then given by (see Ross, 1996; Proposition 8.4.1)

$$\mathbb{E} \left( e^{-\lambda T(i, \phi_\sigma^*(i))} \right) = \int_0^\infty e^{-\lambda T(i, \phi_\sigma^*(i))} f(T(i, \phi_\sigma^*(i))) dT(i, \phi_\sigma^*(i)) \quad (\text{C.2})$$

$$= e^{-\left( \sqrt{(\alpha - \frac{1}{2}\sigma^2)^2 + 2\sigma^2\lambda} - (\alpha - \frac{1}{2}\sigma^2) \right) \frac{\ln \left( \frac{\phi_\sigma^*(i)}{\phi} \right)}{\sigma^2}} \quad (\text{C.3})$$

and can be used to determine the expected waiting time as

$$\mathbb{E}(T(i, \phi_\sigma^*(i))) = \int_0^\infty T(i, \phi_\sigma^*(i)) f(T(i, \phi_\sigma^*(i))) dT(i, \phi_\sigma^*(i)) \quad (\text{C.4})$$

$$= -\lim_{\lambda \rightarrow 0} \frac{\partial \mathbb{E}(e^{-\lambda T(i, \phi_\sigma^*(i))})}{\partial \lambda} = \frac{\ln \left( \frac{\phi_\sigma^*(i)}{\phi} \right)}{\alpha - \frac{1}{2}\sigma^2}. \quad (\text{C.5})$$

<sup>1</sup> A detailed derivation is offered by Karatzas and Shreve (1991, p.196) or by Karlin and Taylor (1975, p.363).

More precisely,

$$\mathbb{E}(T_i(\phi_\sigma^*(i), \phi)) = \begin{cases} \frac{1}{\alpha - \frac{1}{2}\sigma^2} \ln\left(\frac{\phi_\sigma^*(i)}{\phi}\right) & \text{if } \alpha > \frac{1}{2}\sigma^2 \\ \infty & \text{if } \alpha \leq \frac{1}{2}\sigma^2 \end{cases} \quad (\text{C.6})$$

with  $\phi_\sigma^*(i) > \phi$  and  $i \in \{E, F\}$ .

## D Expected Entry Time and Comparative Statics

Exploiting the monotonicity of  $V_0^e$  in  $\phi_\sigma^*(i)$ , we prove that  $\frac{\partial \phi_\sigma^*(i)}{\partial \sigma} > 0$  and  $\frac{\partial \phi_\sigma^*(i)}{\partial \alpha} < 0$  by proving that  $\frac{\partial V_0^e(i, \phi_\sigma^*)}{\partial \sigma} > 0$  and  $\frac{\partial V_0^e(i, \phi_\sigma^*)}{\partial \alpha} < 0$ , respectively.

Rearranging (36), we obtain

$$\frac{M(i)\phi_\sigma^*(i)^\kappa}{\mu_e - \alpha_e} = V_0^e(\phi_\sigma^*(i)) = \frac{\beta_1}{\beta_1 - \kappa} I(i). \quad (\text{D.1})$$

The derivative of  $V_0^e(\phi_\sigma^*(i))$  with respect to  $\sigma$  results as

$$\frac{\partial V_0^e(\phi_\sigma^*(i))}{\partial \sigma} = \frac{\partial \beta_1}{\partial \sigma} I_i \left( \frac{-\kappa}{(\beta_1 - \kappa)^2} \right). \quad (\text{D.2})$$

From B.1 we can derive

$$\frac{\partial \beta_1}{\partial \sigma} = -\frac{\beta_1 \sigma (\beta_1 - 1)}{\sigma^2 (\beta_1 - \frac{1}{2}) + r - (\mu - \alpha)}. \quad (\text{D.3})$$

Substituting into D.2 results in

$$\frac{\partial V_0^e(\phi_\sigma^*(i))}{\partial \sigma} = \frac{V_0^e(\phi_\sigma^*(i)) \sigma (\beta_1 - 1) \kappa}{(\sigma^2 (\beta_1 - \frac{1}{2}) + r - (\mu - \alpha)) (\beta_1 - \kappa)}. \quad (\text{D.4})$$

For  $\beta_1 > 1$  and  $\kappa \geq 1$  the numerator is always positive. We can prove that the denominator is also always positive. To do so, we rewrite (28) as

$$(\beta_1 - \frac{1}{2})\sigma^2 + r - (\mu - \alpha) = \sigma^2 \sqrt{\left( \frac{r - (\mu - \alpha)}{\sigma^2} - \frac{1}{2} \right)^2} + \frac{2r}{\sigma^2} > 0. \quad (\text{D.5})$$

The right-hand side of this equation is always positive for  $\beta_1 > 1$ , and hence  $\frac{\partial V_0^e(\phi_\sigma^*(i))}{\partial \sigma} > 0$ .

Furthermore,

$$\frac{\partial V_0^e(\phi_\sigma^*(i))}{\partial \alpha} = \frac{\partial \beta_1}{\partial \alpha} I_i \left( \frac{-\kappa}{(\beta_1 - \kappa)^2} \right). \quad (\text{D.6})$$

From D.2 we receive

$$\frac{\partial \beta_1}{\partial \alpha} = \frac{-\beta_1}{(\beta_1 - \frac{1}{2})\sigma^2 + r - (\mu - \alpha)} < 0. \quad (\text{D.7})$$

Hence, we have  $\frac{\partial V_i^*}{\partial \alpha} > 0$ . Since  $V_i^*$  behaves as  $\phi_i^*$ , we can state  $\frac{\partial \phi_\sigma^*(i)}{\partial \alpha} < 0 \quad \wedge \quad \frac{\partial \phi_\sigma^*(i)}{\partial \sigma} > 0$ .

With these results we can consider the following partial derivatives of C.6:

$$\frac{\partial \mathbb{E}(T(i, \phi_\sigma^*(i)))}{\partial \sigma} = \frac{\sigma}{(\alpha - \frac{1}{2}\sigma^2)^2} \ln \left( \frac{\phi_\sigma^*(i)}{\phi} \right) + \frac{1}{(\alpha - \frac{1}{2}\sigma^2)} \frac{1}{\phi_\sigma^*(i)} \frac{\partial \phi_\sigma^*(i)}{\partial \sigma} > 0 \quad (\text{D.8})$$

$$\frac{\partial \mathbb{E}(T(i, \phi_\sigma^*(i)))}{\partial \alpha} = -\frac{1}{(\alpha - \frac{1}{2}\sigma^2)^2} \ln \left( \frac{\phi_\sigma^*(i)}{\phi} \right) + \frac{1}{(\alpha - \frac{1}{2}\sigma^2)} \frac{1}{\phi_\sigma^*(i)} \frac{\partial \phi_\sigma^*(i)}{\partial \alpha} < 0. \quad (\text{D.9})$$

In both modes expected entry time increases in  $\sigma$  and decreases in  $\alpha$ .

A longer working paper version of this analysis is published in the discussion papers on Business and Economics No.19/2012 at University of Southern Denmark.